

# Tutorial on Graphical Models and Causal Inference

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# Overview

- Graphical Models
  - undirected graphs
  - directed graphs (DAGs)
- Causality
  - causal graphs
  - intervention graphs
- Identifiability
  - ‘back-door’ / ‘front-door’ criteria
  - lots of examples
- Graphs and counterfactuals
  - Further topics
    - sequential decisions
    - instrumental variables
    - direct / indirect effects

# Graphical Models

Statistical models

- for multivariate rv's  $\mathbf{X} = (X_1, \dots, X_K)$
- characterized through specific conditional independencies
- that have to be representable in a graph.

→ Graph  $\mathcal{G} = (V, E)$  with

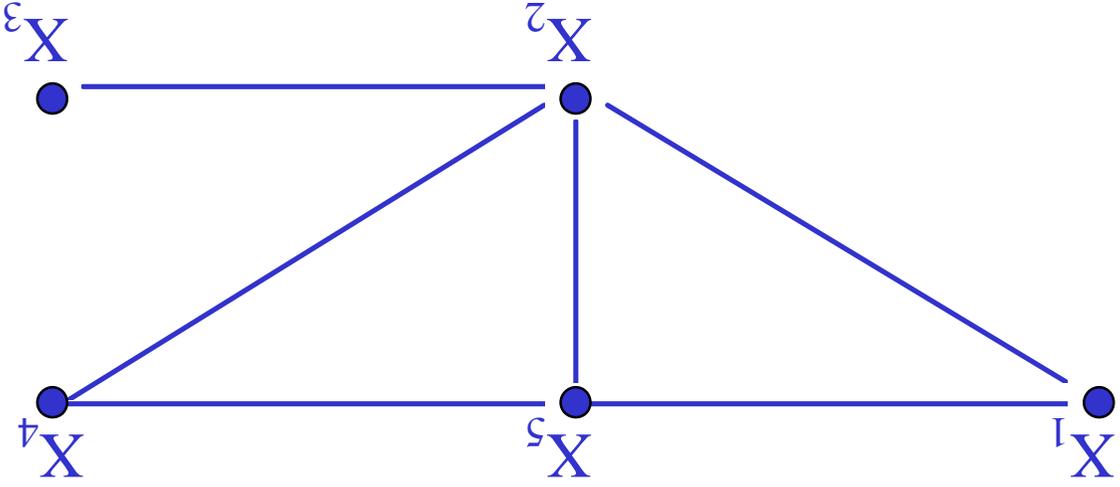
$V = \{1, \dots, K\}$  set of vertices  $\hat{=}$  variables

$E \subset V \times V$  set of edges  $\hat{=}$  conditional dependence

Distinguish:

- Undirected graphs (edges:  $\{i, j\}, i, j \in V$ )
- Directed graphs (edges:  $(i, j)$ )

## Example:



$V = \{1, \dots, 5\}$ ,  $E = \{\{1, 2\}, \{1, 5\}, \{2, 4\}, \{2, 5\}, \{3, 2\}, \{4, 5\}\}$

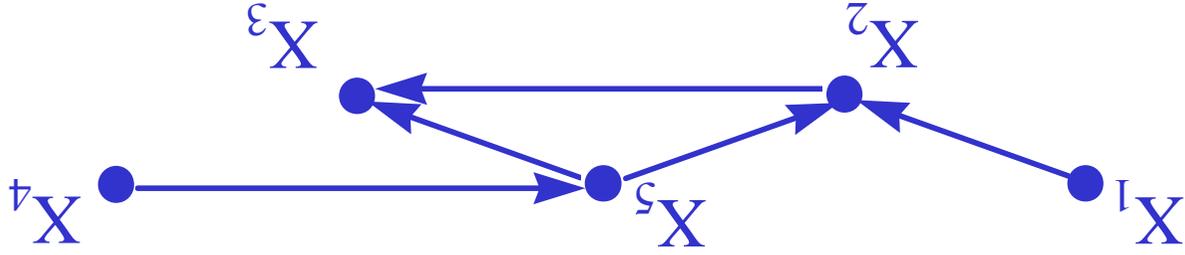
$\Leftrightarrow X_1 \perp\!\!\!\perp X_3 \mid X_2$  and  $X_1 \perp\!\!\!\perp X_4 \mid (X_2, X_5)$  etc.

## Definition:

Distribution  $P_{\mathbf{X}}$  is called  $\mathcal{G}$ -Markov, if for all disjoint  $A, B, C \subset V$ :

$C$  separates  $A$  and  $B$  in  $\mathcal{G} \Leftrightarrow X^A \perp\!\!\!\perp X^B \mid X^C$ .

## Example



$$V = \{1, \dots, 5\}, E = \{(1, 2), (2, 3), (2, 5), (3, 5), (4, 5), (5, 2), (5, 3)\}$$

Notations:  $A \subset V$

$\text{pa}(A) \triangleq$  parents  
 $\text{ch}(A) \triangleq$  children  
 $\text{an}(A) \triangleq$  ancestors  
 $\text{nd}(A) \triangleq$  non-descendants  
 $\text{de}(A) \triangleq$  descendants  
 $\text{An}(A) = \text{an}(A) \cup \{A\}$

**Definition:** Distribution  $P_{\mathbf{X}}$  is called  $\mathcal{G}$ -Markov, if

$$X_i \perp\!\!\!\perp \mathbf{X}_{\text{nd}(i) \setminus \text{pa}(i)} \mid \mathbf{X}_{\text{pa}(i)} \quad \forall i \in V.$$

$\Leftrightarrow$  factorization:  $P(X_1, \dots, X_K) = \prod_{i \in V} P(X_i \mid \mathbf{X}_{\text{pa}(i)})$

$\rightarrow$  system of univariate regressions

Important for the interpretation:

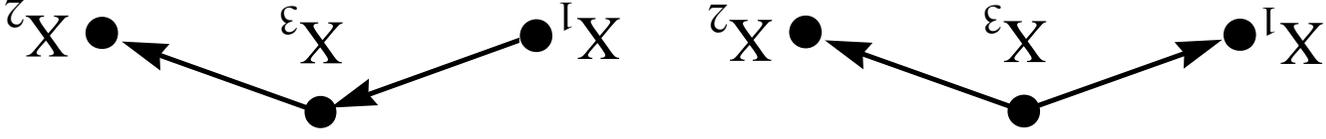
- conditioning on common child (**selection**)  $\Leftrightarrow$  dependence



here:  $X_1 \perp\!\!\!\perp X_2$ , but  $X_1 \not\perp\!\!\!\perp X_2 \mid X_3$

- marginalizing w.r.t. common parent (**confounder**) or intermediate

variables  $\Leftrightarrow$  dependence



here:  $X_1 \perp\!\!\!\perp X_2 \mid X_3$ , but  $X_1 \not\perp\!\!\!\perp X_2$

'Markov equivalence'  $\Leftrightarrow$  DAGs are not causal

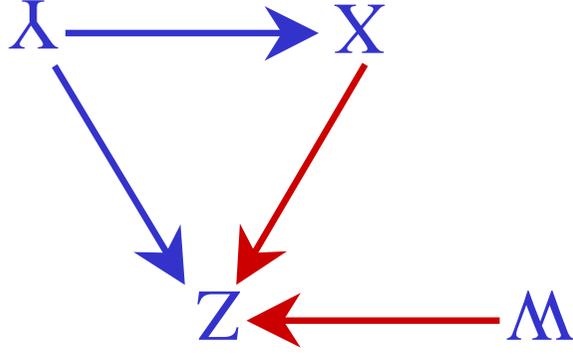
(Verma & Pearl, 1990)

Two DAGs  $G_1$  and  $G_2$  are **Markov equivalent** (encode exactly the same conditional independencies) iff

(i) they have the same 'skeleton'

(ii) they have the same 'V-structures'

$$Z \perp\!\!\!\perp X \text{ and } Z \perp\!\!\!\perp W$$



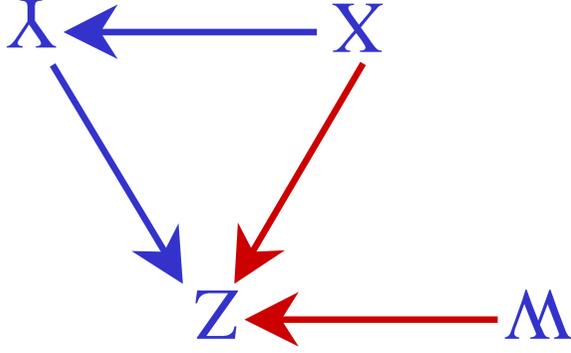
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## Markov equivalence

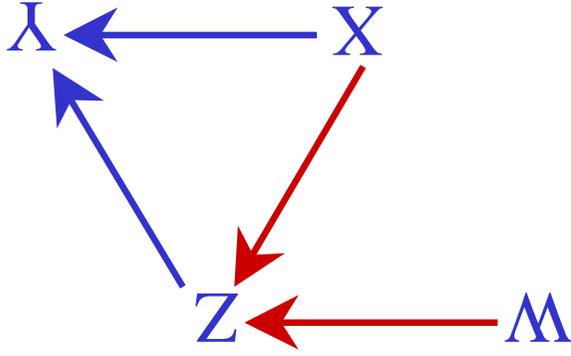
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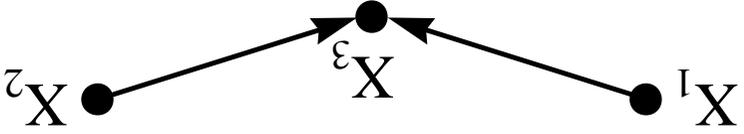
(ii) they have the same 'V-structures'

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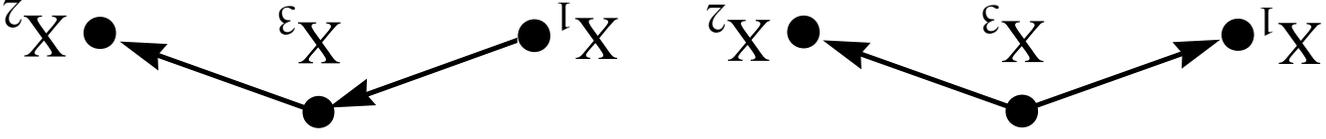
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here:  $X_1 \perp\!\!\!\perp X_2 \mid X_3$ , but  $X_1 \not\perp\!\!\!\perp X_2$

$\Rightarrow$  how do we read off all conditional independencies?

# Separation in DAGs

Construct **moral** graph  $G^m$

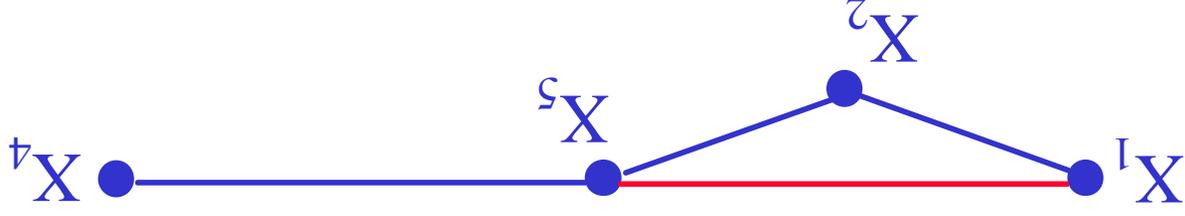
- connect parents of common children
- make all edges undirected

## Theorem

$G$  DAG,  $P_X$   $G$ -Markov, and  $A, B, C \subset V$  disjoint:

$C$  separates  $A$  and  $B$  in  $G^{An(A \cup B \cup C)}$   $\Leftrightarrow X^A \perp\!\!\!\perp X^B \mid X^C$

**Example** (continued):  $G^{An(2,4,5)}$



$\Leftrightarrow X^2 \perp\!\!\!\perp X^4 \mid X^5$

## Alternatively: d-Separation

Given DAG  $\mathcal{G} = (V, E)$ . A path between  $X$  and  $Y \in V$  is **blocked** by  $S \subset V \setminus \{X, Y\}$  if

(i) it **contains** a constellation  $\rightarrow Z \rightarrow$  or  $\rightarrow Z \rightarrow$  and  $Z \in S$  or

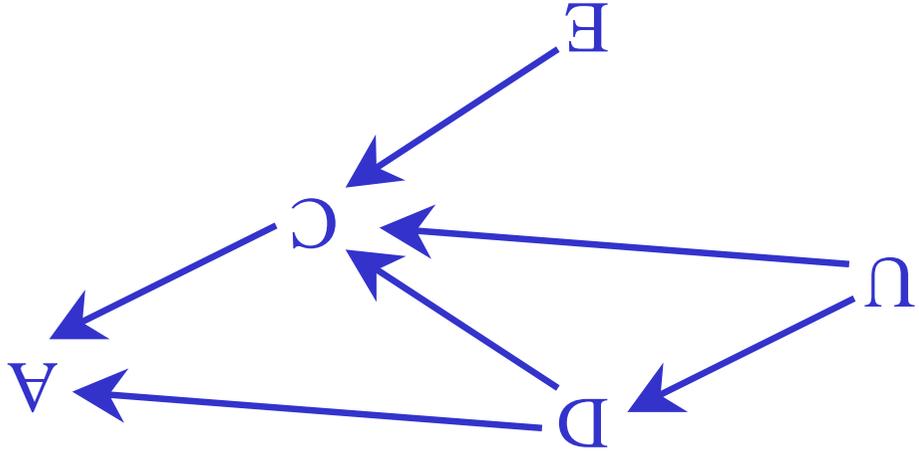
(ii) it **contains** a constellation  $\rightarrow Z \rightarrow$  and **neither**  $Z$  nor any **descendants of**  $Z$  are elements of  $S$ .

$A$  and  $B \subset V$  are d-separated by  $S \subset V \setminus (A \cup B)$  if every path between  $A$  and  $B$  is blocked by  $S$ .

$\Leftrightarrow$  separation using moralisation criterion, i.e.  $\text{ALL} B \mid S$

## Example

$D$  = endometrial cancer  
 $E$  = estrogens  
 $C$  = vaginal bleeding  
 $A$  = ascertained cancer  
 $U$  = unknown uterine abnormality



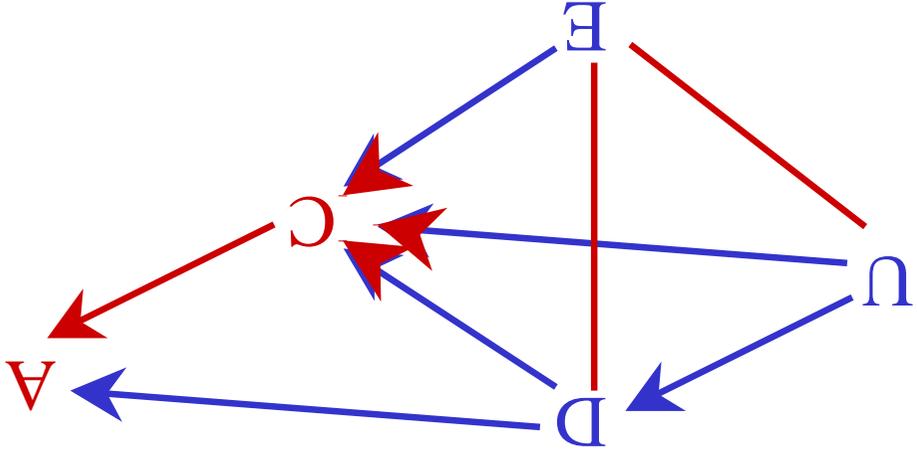
Graph implies:  $E \perp\!\!\!\perp (D, U)$  and  $A \perp\!\!\!\perp (U, E) \mid (D, C)$

- $H_0$ : no 'causal effect' of  $E$  on  $D$  implies no association
- Case control study

(Robins, 2001)

## Example

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- $H_0$ : no 'causal effect' of  $E$  on  $D$  implies no association
- Case control study  $\Rightarrow$  **always conditional on  $A$**

$\Rightarrow H_0$  cannot be tested in this design as  $E \perp\!\!\!\perp D \mid A$  and  $E \perp\!\!\!\perp D \mid (A, C)$

# Causality: A Basic Distinction

**seeing**  $\leftrightarrow$  **doing**

**Association:** **seeing**  $X$  helps to predict  $Y$

$\Rightarrow$  usual conditional probability  $P(Y | X = x)$

**Causation:** **setting**  $X$  helps to predict  $Y$

$\Rightarrow$  new notation, e.g.

$P(Y | \text{do}(X = x))$  or  $P(Y || X = x)$  or

$P(Y | X, F_X = x)$  where  $F_X$  **intervention** indicator

# Causal Graphs

(Lauritzen, 2000)

**Definition:** DAG  $\mathcal{G}$ ,  $P_{\mathbf{X}}$   $\mathcal{G}$ -Markov. Then,  $\mathcal{G}$  causal wrt  $B \subset V$  if for any  $A \subset B$

$$P(\mathbf{X} \mid \text{do}(\mathbf{X}_A = a)) = \prod_{i \in V \setminus A} P(X_i \mid \mathbf{X}_{\text{pa}(i)}) \Big|_{\mathbf{X}_A = a}$$

in words:

- all conditional specifications on  $V \setminus A$  remain the same
- $\mathbf{X}_A$  are simply fixed to  $a$  when appearing in  $\mathbf{X}_{\text{pa}(i)}$

## Example 1



- as DAG:  $P(A, B) = P(B)P(A | B)$  — **no restrictions**

- if **causal wrt B**:  $P(A | \text{do}(B = b)) = P(A | B = b)$

- if **not causal** (association due to other factor):  $P(A | \text{do}(B = b)) = P(A)$

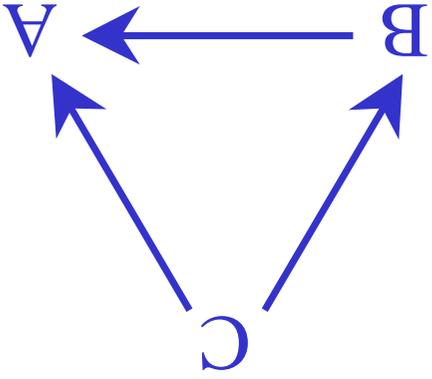
NOTE: in causal graph  $\text{do}(B = b)$  is like usual conditioning if  $\text{pa}(B) = \emptyset$ .

## Example 2

- as DAG:  $P(A, B, C) = P(C)P(B | C)P(A | B, C)$
- if causal wrt  $B$ :

$$P(A, C | \text{do}(B = b)) = P(C)P(A | B = b, C)$$

→ remove arrow into  $B$  and condition on  $B$



**Causal effect of  $B$  on  $A$  is**

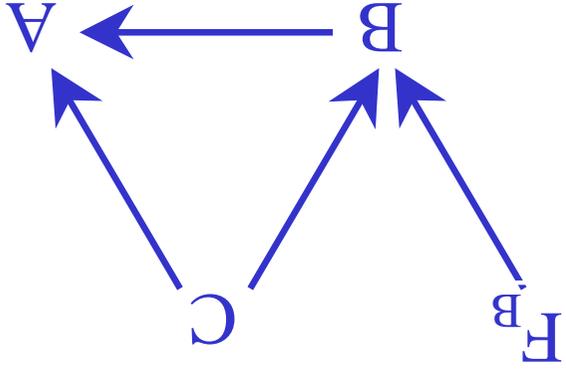
$$P(A | \text{do}(B = b)) = \sum_C P(C)P(A | B = b, C)$$

which is **not** the usual conditioning

$$P(A | B = b) = \frac{\sum_C P(C)P(B | C)}{\sum_C P(C)P(A | B = b, C)}$$

# Alternative: Influence Diagrams

(David, 2002)



Define **intervention indicator**

$$F_B = \begin{cases} \emptyset & = \text{idle} = \text{no intervention} \\ b & = \text{active} = \text{do}(B = b) \end{cases}$$

so that  $P(B | F_B = \emptyset, C) = \text{observational distribution}$

and  $P(B | F_B = b, C) = I_{\{B=b\}}$  when **intervening**

- intervention is made explicit in the graph
- graph is again 'just' conditional independence graph
- can be generalised to random / conditional interventions
- NOTE: statements always conditional on  $F_B$

**Example:**

- DAG implies

$do_{F_B}$  and  $do_{F_B}(B, C)$  or

$$P(A, B, C | F_B) = P(C)P(B | F_B, C)P(A | B, C)$$

- under intervention

$$P(A, C | F_B = b) = P(C)I_{\{B=b\}}P(A | B = b, C) \rightarrow \text{same as before}$$

**Causal effect**

$$P(A | F_B = b) = \sum_C P(C)P(A | B = b, C)$$

# Use of Causal Graphs

→ simple and intuitive (?) 'language' for complex multivariate dependence structures

→ important characteristics of the statistical model can be read off the graph, e.g. independencies under selection (case control studies) or latent structures (marginalisation)

## Identifiability

- can we identify causal effect from observational data?
- what do we need to condition on?
- what must we not condition on?

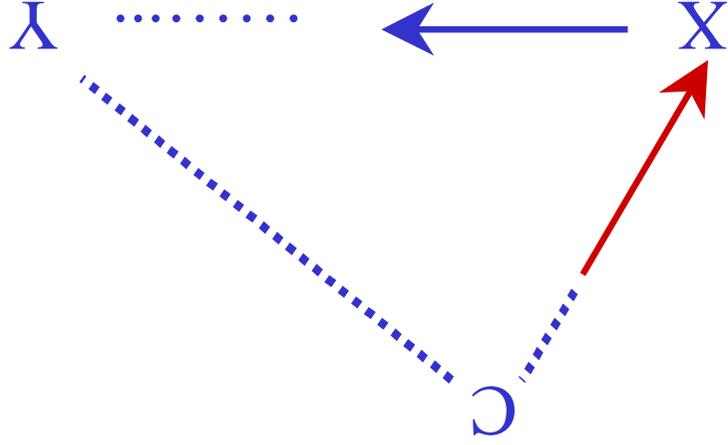
→  $P(Y | \text{do}(X = x))$  or  $P(Y | F_X = \text{'active'})$  has to be expressed with observable quantities only.

# 'Back-Door' Criterion with Causal Graphs

## Theorem

Given causal graph  $\mathcal{G}$  on  $V$ . Then  $C \subset V \setminus \{X, Y\}$  identifies causal effect of  $X$  on  $Y$  if

- (i)  $C$  is non-descendant of  $X$  and
- (ii) all 'back-door' paths from  $X$  to  $Y$  are blocked by  $C$



# 'Back-Door' Criterion with Intervention Graphs

## Theorem

Given intervention DAG  $\mathcal{G}$  on  $V$  with  $X, Y, F_X \in V$ . Then  $C \subset V \setminus \{X, Y, F_X\}$  identifies causal effect of  $X$  on  $Y$  if

(!)  $C \perp\!\!\!\perp F_X$  ( $\Leftrightarrow C$  non-descendant of  $X$ )

(!!)  $Y \perp\!\!\!\perp F_X \mid (X, C)$  ( $\Leftrightarrow C$  blocks back-door paths)

Causal effect is then calculated as

$$P(Y \mid F_X = x) = \sum_C P(C)P(Y \mid X = x, C)$$

## How can we use the Back-door Criterion?

- Construct the DAG based on knowledge of
  - subject matter (basic biology etc.)
  - temporal ordering
  - study design
  - statistical evidence
- ⇒ will typically include unobservable variables
- check for which choice of  $C$  (if any) (i) and (ii) hold
  - check for separations

# Example: Drug Use During Pregnancy

(Robins, 2001)

Exposure =  $E$  = drug intake during pregnancy (from medical record)  
 Disease =  $D$  = congenital defect of child

Surrogate =  $E^*$  = self reported drug use (after diagnosis of  $D$ )

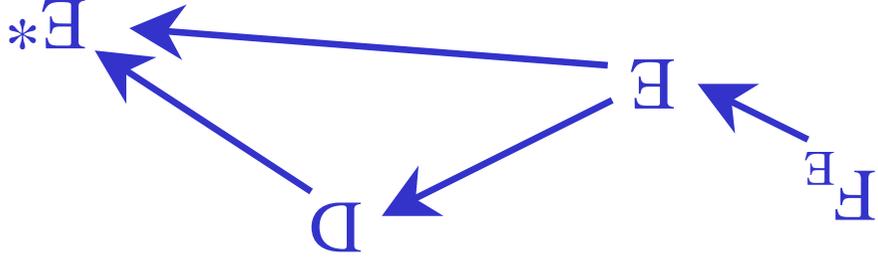
→ case-control study

$$D \perp\!\!\!\perp E \mid E$$

⇒ causal effect of  $E$  on  $D$

can be identified through *OR<sub>ED</sub>*

$$F_E \not\perp\!\!\!\perp E^*$$



NOTE: should and **must** ignore  $E^*$ !

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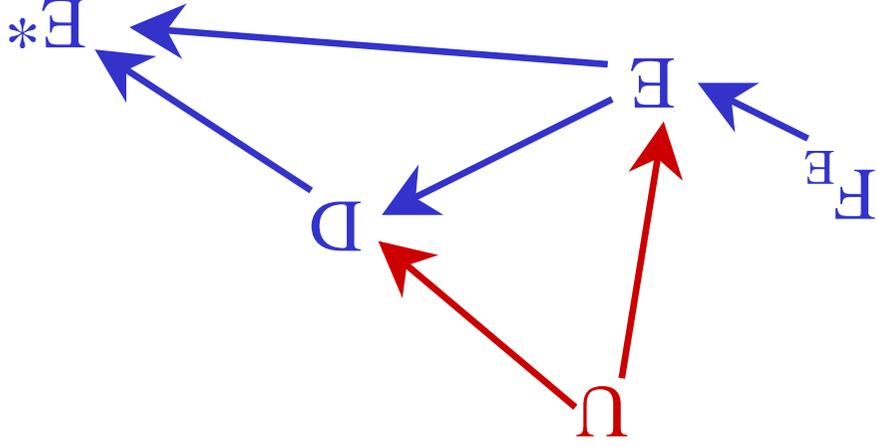
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$$D \not\perp F_E | E$$

⇒ causal effect of  $E$  on  $D$   
 can **not** be identified



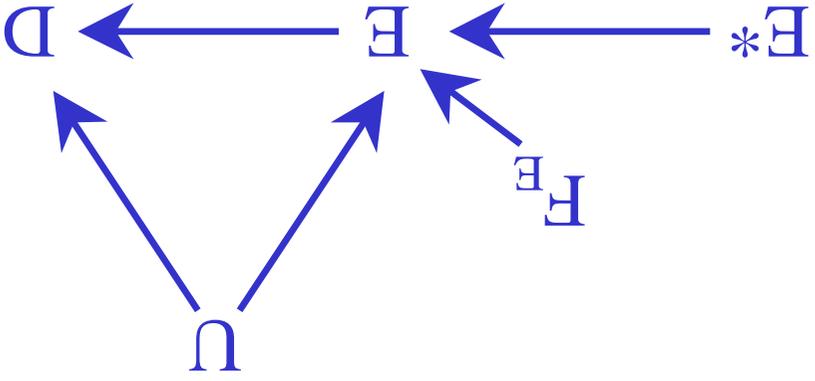
# Example: Partial Compliance

$E$  = actual drug intake

$D$  = recovery

$E^*$  = assigned drug intake

→ randomized trial



randomisation  $\Rightarrow E^* \perp U$

and further  $D \perp\!\!\!\perp E^* \mid (E, U)$

but  $D \not\perp\!\!\!\perp F_E \mid E$  and  $D \not\perp\!\!\!\perp F_E \mid (E, E^*)$

$\Rightarrow$  causal effect can **not** be identified

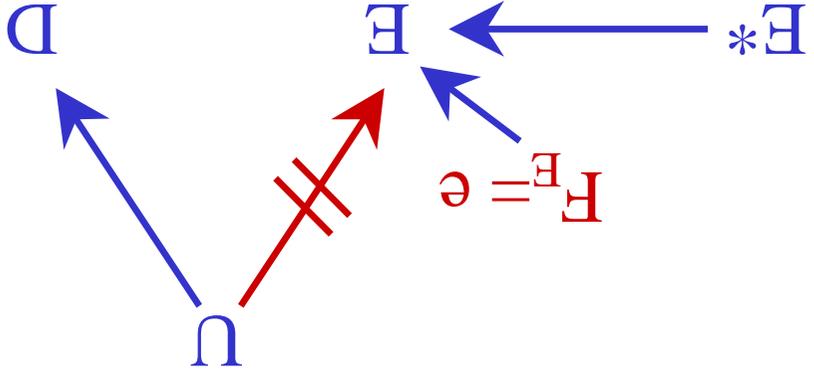
## Example: Partial Compliance

$E$  = actual drug intake

$D$  = recovery

$E^*$  = assigned drug intake

$\rightarrow$  randomized trial



$H_0$  : no causal effect of  $E$  on  $D$   $\Rightarrow E^* \perp\!\!\!\perp D$

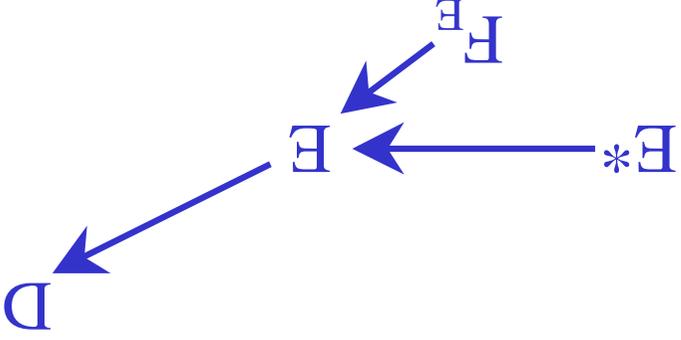
$\Rightarrow H_0$  can be tested!

NOTE: can compute bounds for causal effect of  $E$  on  $D$ .

# Example: Mineworkers Exposed to Radon

(Robins, 2001)

$E$  = exposure to radon (lung dosimetry)  
 $D$  = mortality (no censoring)  
 $E^*$  = radon air level in mine  
→ prospective cohort study  
data shows:  $E^*$  and  $D$  associated given  $E$



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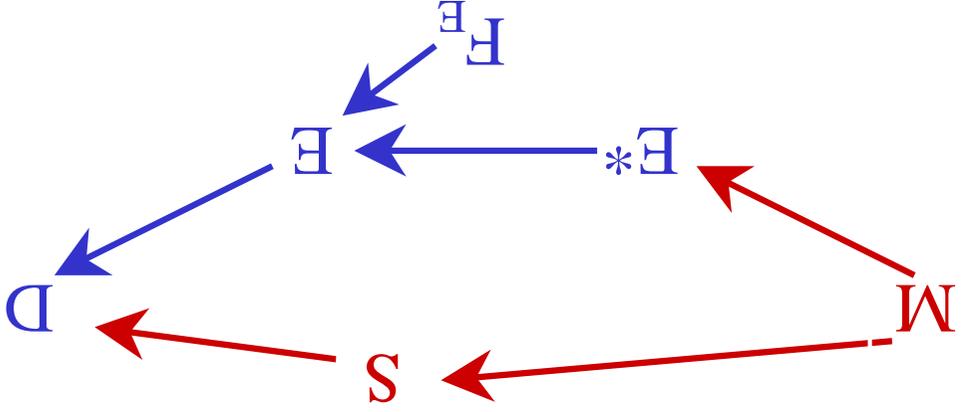
→ prospective cohort study

data shows:  $E^*$  and  $D$  associated given  $E$

$$D \not\perp E^* | E$$

$$\text{but } D \perp\!\!\!\perp E^* | (E, E^*) \text{ and } E^* \perp\!\!\!\perp E$$

$\Rightarrow E^*$  identifies the causal effect



$S$  = silica

$M$  = mine

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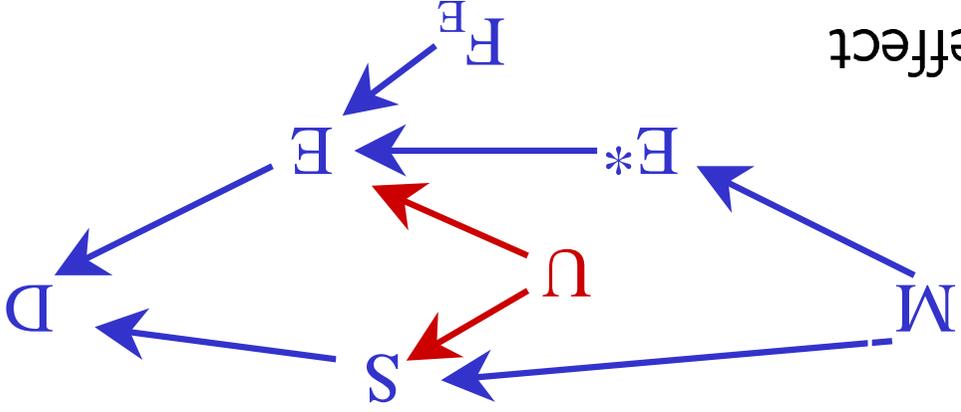
→ prospective cohort study

data shows:  $E^*$  and  $D$  associated given  $E$

e.g.  $U$  = physical exertion

now  $D \not\perp\!\!\!\perp F_E \mid (E, E^*)$

$\Rightarrow E^*$  does not identify the causal effect



$S$  = silica

$M$  = mine

## Front-door criterion

### Theorem

Given intervention DAG  $g$  on  $V$  with  $X, Y, F_X \in V$ . Let  $C \subset V \setminus \{X, Y, F_X\}$  and  $U = V \setminus \{X, Y, F_X, C\}$ . Then  $C$  identifies causal effect of  $X$  on  $Y$  if

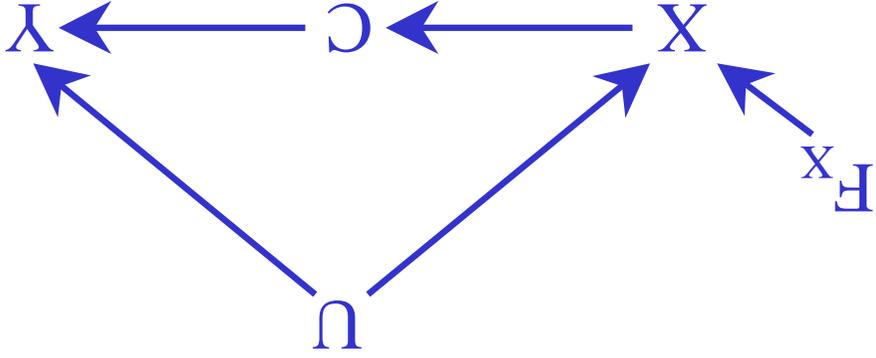
$$(I) \quad U \perp\!\!\!\perp F_X$$

$$(II) \quad C \perp\!\!\!\perp (U, F_X) \mid X$$

$$(III) \quad Y \perp\!\!\!\perp F_X \mid (C, U)$$

Causal effect is then calculated as

$$P(Y \mid \text{do}(X = x_*)) = \sum_C P(C \mid X = x_*) \sum_X P(Y \mid C, X) P(X)$$



# Graphs and Counterfactuals

consider binary treatment  $X \in \{0, 1\}$

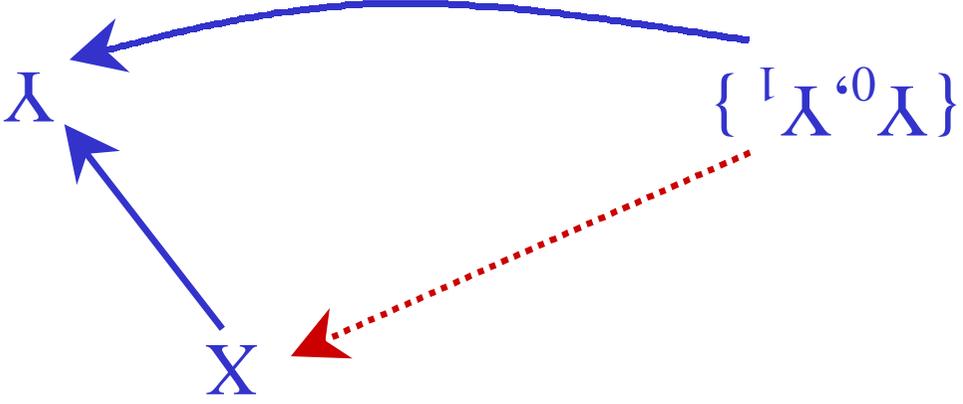
$Y_0 =$  response if  $X = 0$

$Y_1 =$  response if  $X = 1$  for **same** subject (at the same time)

$\Rightarrow$  can **never** be observed together

$\{Y_0, Y_1\}$  unobserved variable **always** affecting  $Y$

$\Rightarrow X$  must **not depend** on  $\{Y_0, Y_1\}$



# Graphs and Counterfactuals

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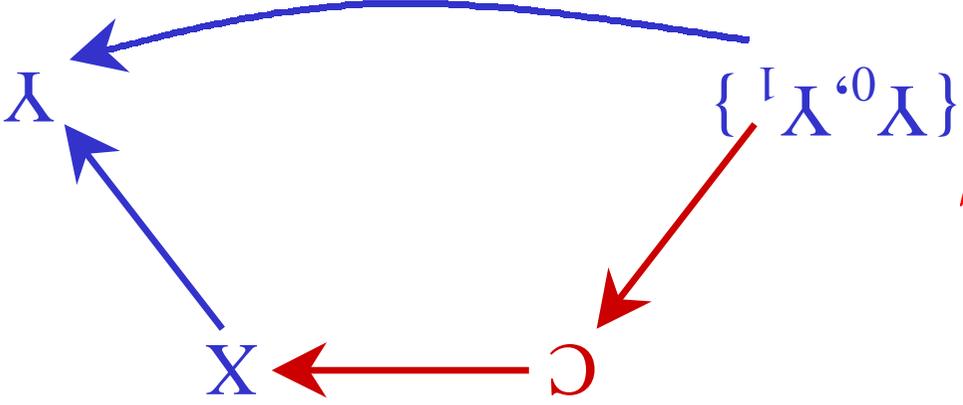
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$\{Y_0, Y_1\}$  unobserved variable **always** affecting  $Y$

$\Rightarrow X$  must **not** depend on  $\{Y_0, Y_1\}$

possibly **conditional** on covariates  $C$



# Counterfactuals, Graphs and Structural Equations

(Pearl, 2000)

$X$  = treatment / exposure,  $Y$  = response,  $C$  = covariate

**Structural equation model (SEM)** — ingredients:

• DAG  $\mathcal{G}$

• **equations**:  $X := f_X(\text{pa}(X), U_X)$

$Y := f_Y(\text{pa}(Y), U_Y)$

$C := f_C(\text{pa}(C), U_C)$

where  $f_X, f_Y, f_C$  describe 'stable' functional relations

• probability distribution on  $(U_X, U_Y, U_C)$

$\Leftrightarrow$  induce probability distribution on  $(X, Y, C)$

which is  $\mathcal{G}$ -Markov if  $U_X, U_Y, U_C$  independent

in this framework we have

$$Y_0 = f_Y(\text{pa}(Y) \setminus X, X = 0, U_Y)$$

$$Y_1 = f_Y(\text{pa}(Y) \setminus X, X = 1, U_Y)$$

with the same  $U_Y$

$\Leftrightarrow$  distribution on  $(U_X, U_Y, U_C)$  also induces a probability distribution on  $(Y_0, Y_1, X, C)$

... in particular a **joint distribution for  $(Y_0, Y_1)$** !

$\Leftrightarrow$  allows to address questions like

“Would Mr. Smith (who has cancer) have developed cancer had he not (counter to the fact) been exposed to radiation?”

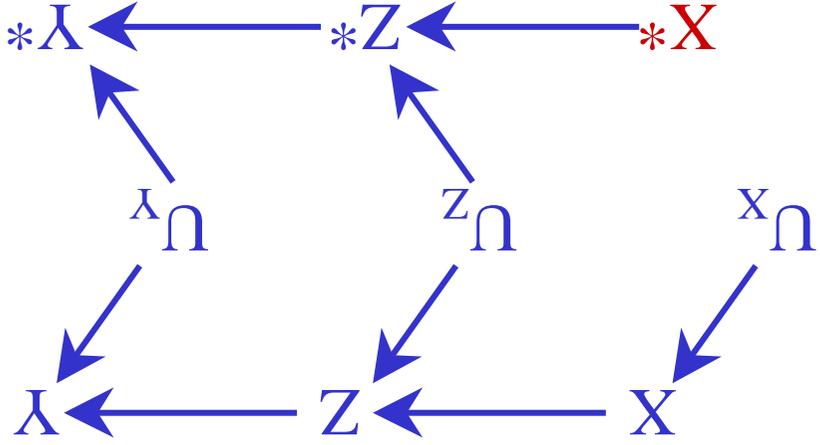
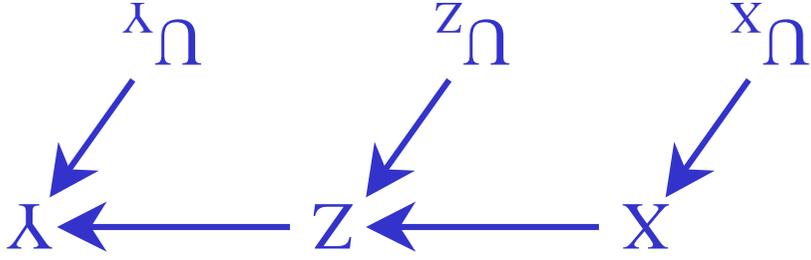
$$\text{e.g. } P(Y_0 = 1 \mid X = 1, Y = 1)$$

# Twin Networks

(Balke & Pearl, 1994)

**Idea:** include counterfactual variables in graph such that conditional independencies of joint distribution are represented.

**Example**



observational graph according to SEM

graph with counterfactuals when **setting  $X^*$**  read off, e.g.  $Z \perp\!\!\!\perp X^* \mid Y$

## Further Topics

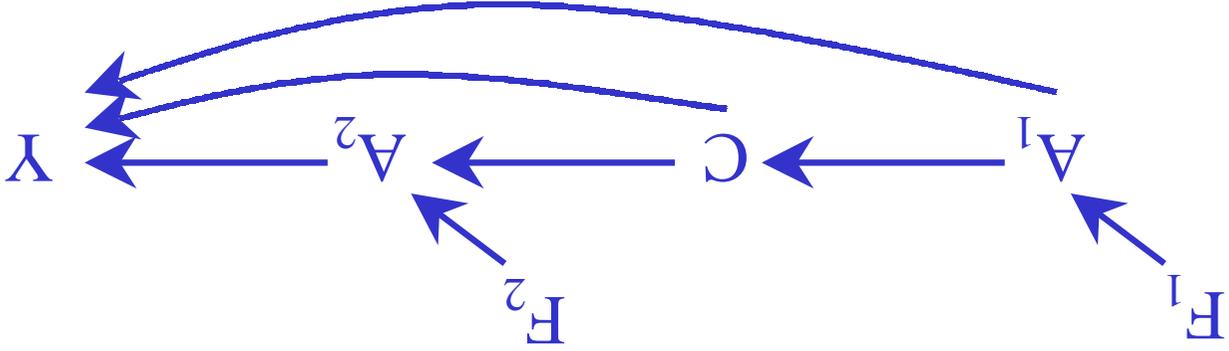
- Sequential decisions / G-formula
- Instrumental variables
- Direct / indirect effects

# Sequential Decisions / G-Formula

DAG implies

$Y \perp\!\!\!\perp F_1 \mid A_1$  and

$Y \perp\!\!\!\perp F_2 \mid (A_1, A_2, C)$  and  $C \perp\!\!\!\perp F_1 \mid A_1$



or  $P(Y, A_1, A_2, C \mid F_1, F_2) = P(A_1 \mid F_1) P(C \mid A_1) P(A_2 \mid F_2, A_1, C) P(Y \mid A_1, A_2, C)$

$\Rightarrow$  Causal effect (G-Formula)

(Robins, 1986)

$$P(Y \mid F_1 = a, F_2 = b) = \sum_C P(C \mid A_1 = a) P(Y \mid C, A_1 = a, A_2 = b)$$

NOTE:  $P(Y \mid F_1 = a, F_2 = b) \neq P(Y \mid A_1 = a, A_2 = b)$

# Instrumental Variable

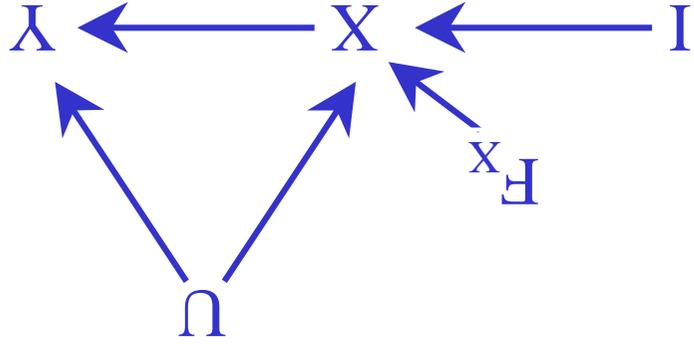
(Pearl, 1995)

**Problem:** causal effect of  $X$  on  $Y$  confounded by unobserved  $U$ .

$\Rightarrow$  an **instrumental variable**  $I$  satisfies:

(!)  $I \perp\!\!\!\perp Y \mid (X, F_X = x)$  and

(!!)  $I \not\perp\!\!\!\perp X$



$\Rightarrow$  allows to calculate bounds for the causal effect

# Instrumental Variable

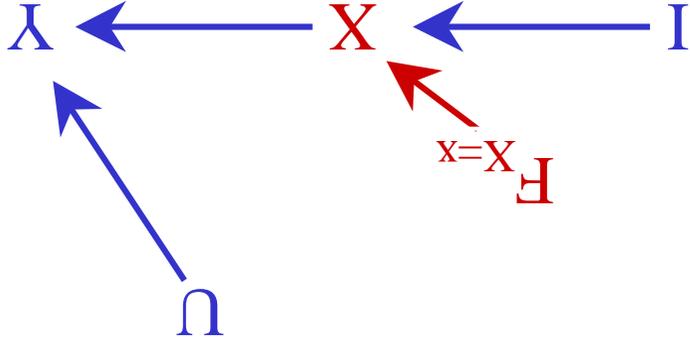
(Pearl, 1995)

**Problem:** causal effect of  $X$  on  $Y$  confounded by unobserved  $U$ .

$\Rightarrow$  an **instrumental variable**  $I$  satisfies:

(!)  $I \perp\!\!\!\perp Y \mid (X, F_{X=x})$  and

(!!)  $I \not\perp\!\!\!\perp X$



$\Rightarrow$  allows to calculate bounds for the causal effect

NOTE: (!) can not easily be tested as typically  $I \perp\!\!\!\perp Y$  and  $I \not\perp\!\!\!\perp X$

# Instrumental Variable

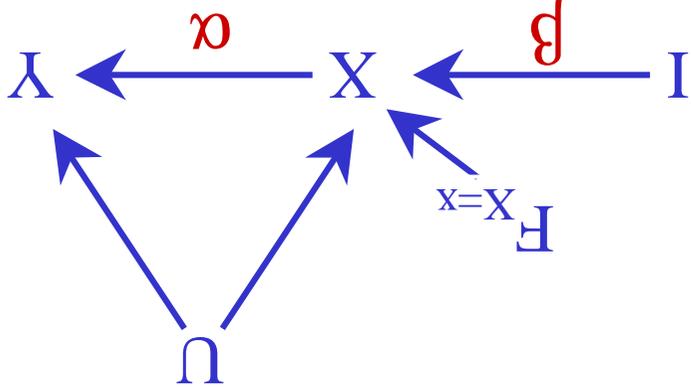
(Pearl, 1995)

**Problem:** causal effect of  $X$  on  $Y$  confounded by unobserved  $U$ .

$\Rightarrow$  an **instrumental variable**  $I$  satisfies:

(!)  $I \perp\!\!\!\perp Y \mid (X, F_{X=x})$  and

(!!)  $I \not\perp\!\!\!\perp X$



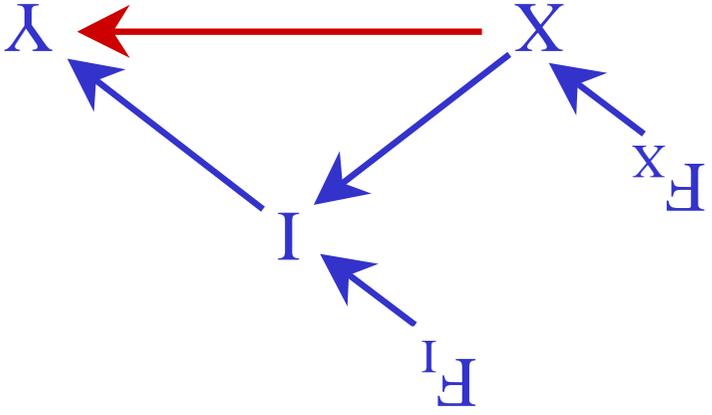
$\Rightarrow$  allows to calculate the causal effect exactly in **linear** (multivariate normal) case:

$$\beta = r_{IX} \quad \text{and} \quad \alpha\beta = r_{IY} \quad \Leftrightarrow \quad \frac{r_{IX}}{r_{IY}} = \alpha$$

where  $r^{ab}$  LS regression coefficients

## Direct Effect

$X$  = treatment e.g. contraceptive pill  
 $I$  = intermediate variable e.g. pregnancy  
 $Y$  = response e.g. thrombosis

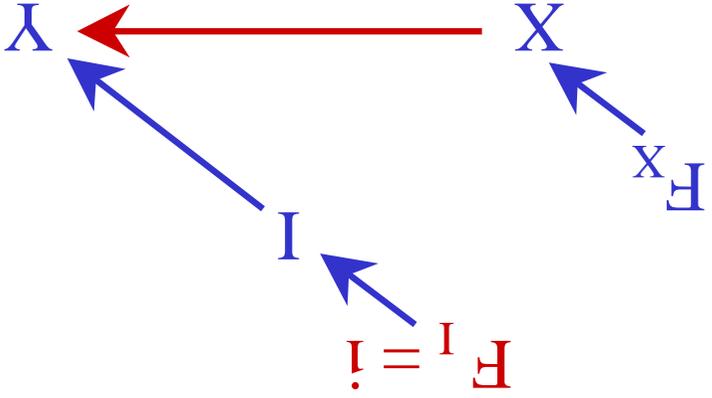


## Direct Effect

$X$  = treatment e.g. contraceptive pill

$I$  = intermediate variable e.g. pregnancy

$Y$  = response e.g. thrombosis



## Direct effect

= effect of setting  $X$  while holding  $I = ?$  constant

=  $P(Y|F_X = x, F_I = ?)$  compared to  $P(Y|F_X = x', F_I = ?)$

(cf. 'controlled direct effect' in Pearl, 2001)

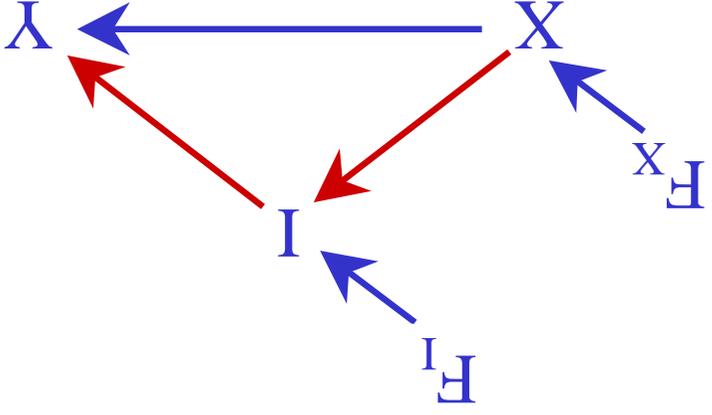
$\{F_X = x\}$  = taking the contraceptive pill

$\{F_X = x'\}$  = not taking the contraceptive pill

$\{F_I = ?\}$  = preventing pregnancy by other method

# Indirect Effect: The Placebo Effect

$X$  = treatment, active agent or not  
 $I$  = psychol. effect of receiving a pill or not  
 $Y$  = response, e.g. recovery

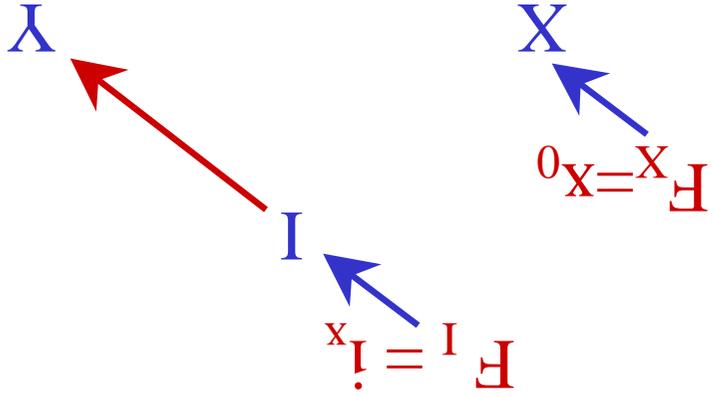


# Indirect Effect: The Placebo Effect

$X$  = treatment, active agent or not

$I$  = psychol. effect of receiving a pill or not

$Y$  = response, e.g. recovery



## Indirect effect

= effect of setting  $I$  to values that arise for different  $X$  while  $X = x_0$  is actually held constant such that  $x_0$  has no direct effect  
 $= P(Y|F_X = x_0, F_I = I_x)$  compared to  $P(Y|F_X = x_0, F_I = I'_x)$

$\{F_X = x_0\}$  = no active agent given

$\{F_I = I_x\}$  = no pill given

$\{F_I = I'_x\}$  = placebo given

# Summary

## Graphs help

- to make explicit the available / prior knowledge
- to make explicit the assumptions required for causal inference
- and to compare the two

## Caution

- eliciting prior knowledge and formulating a graph is not easy
- all results provided are 'non-parametric'
- back-door / front-door criteria are sufficient, not necessary
- DAGs difficult to combine with events in continuous time / feedback systems

# References

## Introductory / overview texts

- Edwards D (2000): Causal inference. Chapter 8 in *Introduction to graphical modelling*. Springer, New York.
- Lauritzen SL (2000): Causal inference from graphical models. In *Complex Stochastic Systems*, Eds. OE Bardorf-Nielsen, DR Cox, C Klüppelberg, pp. 63-107. CRC Press, London.
- Pearl J (2000): *Causality – models, reasoning and inference*. Cambridge University Press.
- Pearl J (2003): Statistics and causal inference: a review. *Test*, 12, 2, pp. 281-345.

## Further references

- Balke AA & Pearl J (1994): Counterfactual probabilities: computational methods, bounds and applications. In *Proceedings of the 10th Conference on Uncertainty in Artificial Intelligence*, pp. 46-54.
- David AP (2002): Influence diagrams for causal modelling and inference. *Int. Stat. Review*, 70, 2, pp. 161-189
- Pearl J (1988): Probabilistic reasoning in intelligent systems. Morgan Kaufmann, CA.

Pearl J (1995): Causal diagrams for empirical research (with discussion). *Biometrika*, 82, 669-710.

Pearl J (1995): Causal inference from indirect experiments. *Artificial intelligence in Medicine*, 7, pp. 561-582.

Pearl J (2001): Direct and Indirect Effects. In *Proceedings of the 17th Conference on Uncertainty in Artificial Intelligence*, pp. 411-420.

Robins JM (1986): A new approach to causal inference in mortality studies with sustained exposure periods — application to control of the healthy worker effect. *Mathematical Modelling*, 7, pp. 1393-1512.

Robins JM (2001): Data, design and background knowledge in etiologic inference. *Epidemiology*, 12, 3, pp. 313-320.

Verma T & Pearl J (1990): Equivalence and synthesis of causal models. In *Proceedings of the 6th Conference on Uncertainty in Artificial Intelligence*, pp. 220-227.