# DEVELOPMENT OF FINANCIAL MARKETS AND ECONOMIC GROWTH: POLAND AGAINST THE BACKGROUND OF SELECTED EU COUNTRIES

### Abstract

The impact of financial market development on economic growth is one of the controversial problems in the theory of economics. Results of the R. Levine and other economists' research seem to confirm the statistically significant relationships between financial market development and economic growth. The Polish financial market has a relatively short history (since 1990).

The aim of the paper is to answer the question if there was an impact of the financial market development on economic growth in Poland, Greece, Ireland and Italy in the years 1994-2007 and, if yes, how strong it was. The paper outlines also selected methods and results of other authors' research The author has presented results of his own research based on the econometric analyses of available data concerning economies of the above mentioned countries.

Key words: financial growth, financial markets, bonds, stocks, OLS, capital, GDP JEL CODE: G0, G1

### Introduction

The research which has resulted in this paper aimed at answering the following questions:

1) What are the relationships between financial development and economic growth from the point of view of the theory of economics and finance?

2) Does the development of the financial market in Poland stimulate economic growth in comparison with selected EU countries and, if yes, in what way?

3) How strong is the influence of the financial market development on economic growth in Poland in comparison with selected EU countries?

The research was conducted in three stages. The first stage embraced literature review in the field of relationships between the financial market development and economic growth in the theory of economics from the point of view of results of empirical studies conducted by different authors. This review allowed the author to formulate the basic stylized facts already discovered in economics. What is more, the research methods and techniques applied in the

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research into correlations between financial market development and economic growth were reviewed and analyzed. This enabled the author to define and choose a research method and model used in the empirical studies of the financial market impact on economic growth in Poland and selected EU countries.

The second stage consisted of statistical data collecting and selecting and an analysis of the main trends in the financial market development in Poland and selected EU countries.

The third stage included an analysis of relationships between financial market development and economic growth in Poland and selected EU countries in the years 1994-2007 on the basis of the financial development indicators possible to use. This research took advantage of a simple multi-equation model taking into account conclusions drawn from the methodology of examining this type of relationship applied by R.G. King and R. Levine.

# 1. Financial market development and economic growth – a review of theoretical aspects

Financial market (including monetary market, capital market, credit-deposit market, currency market and derivative instrument market) is one of the key markets in economy related to the market of products and services and labour market through demand-supply and price coupling. On the one hand, it is a plane on which short-, medium- and long-term financial transactions are accomplished and, on the other hand, it is the mechanism of short- and long-term capital mobilisation and its allocation in order to finance investment ventures .

Financial market performs crucial functions in economy, such as: ensuring liquidity in economy, appropriation of financial claims, mobilisation of part of national income above current expenditure for investment purposes, strengthening of the motivating role of profit, allocation of capital in economy, economic shock absorption through the risk sharing mechanism.

It is also noteworthy that the role of financial markets in economy depends on the financial structure of economy. The financial structure of a given country's economy is shaped by: institutions, financial technology, rules of the game, which determine how financial activities are organised at a point in time. R. Stulz uses an analogy that ,, (...) financial structure is to financial system what a foundation is to a house. Many different houses can be built on the same foundation. However, at the same time, a foundation makes it impossible to build some types of houses. If the foundation is designed for a one floor house, it cannot be used to build a skyscraper<sup>1</sup>. Thus, the financial systems functioning in different countries can be divided

<sup>&</sup>lt;sup>1</sup> Stulz (2004, p. 146-147)

into two categories: market-oriented financial systems, referred to as Anglo-Saxon ones, and bank-oriented financial systems, also known as the continental ones. The former ones and their segments are strongly competitive in relationship to the banking sector as an alternative to capital allocation and raising capital. Financial market plays the main role in capital allocation. Analytical companies connected with the financial market provide information to the entire market. The financial market facilitates complex risk management as the signals received from the market allow investors to assess risk and profitability of investments and enterprises. It also facilitates takeovers and mergers of enterprises which, on the one hand, lead to capital concentration and, on the other hand, exert pressure on managers to work effectively and achieve high profitability of enterprises and investment projects.<sup>2</sup> In the bank-oriented system, banks play the main role. They collect information about enterprises and managers and on the basis of an analysis they allocate capital, enable management of different types of risk and in this way affect effectiveness of investment projects in economy. E.R. Siri and P. Tufano point out the role of banks in capital mobilisation in order to finance ventures leading to attainment of the benefit of scale.<sup>3</sup>

Several important factors affect the shape of the financial system model. The level of economic development is the first one. Countries of growing GDP per capita tend to evolve towards a more distinct role of the financial market. An important role is played by a tendency to risk on the side of economic entities (enterprises and households) of the system which is culturally conditioned. The choice of particular forms of financing by enterprises significantly affects the development of the financial market and/or the banking system. One cannot ignore effectiveness of the legal system which regulates financial market functioning and protects shareholders.<sup>4</sup>

It should be emphasized that at present in many countries financial systems combine, to a different degree, the elements of both models. The United States, Great Britain, Canada, Switzerland represent some of the countries where financial systems are definitely market-oriented and where the ratio of stock-exchange capitalisation to GDP is higher than the ratio of bank credit to GDP. However, the second model in which the above mentioned ratio is reverse, is dominant, among others, in Japan, South Korea, Germany and Austria.<sup>5</sup>

<sup>&</sup>lt;sup>2</sup> See: Allen, Gale (2000); (1991).

<sup>&</sup>lt;sup>3</sup> Siri, Tufano (1995, p. 81-128).

<sup>&</sup>lt;sup>4</sup> See also: Osiński, et al. (2004, p. 15).

<sup>&</sup>lt;sup>5</sup> (see; ibidem p. 16).

J. Schumpeter,<sup>6</sup> pointed out specific functions of financial intermediation and financial markets, essential for economic growth and development and consisting in mobilisation of savings, capital allocation, risk management, facilitating transactions, and company monitoring. Taking into account Schumpeter's theory of entrepreneur and innovation one can propose a thesis that also in the case of financial institutions and financial intermediation a process of "creative destruction" occurs which results in financial development being a component of economic development. R. Levine referred to J. Schumpeter's concept. According to him "Financial development occurs when financial instruments, markets, and intermediaries ameliorate – though do not necessarily eliminate – the effects of information, enforcement, and transactions costs and therefore do a correspondingly better job at providing the five financial functions:

- production of ex ante information about possible investments and capital allocation,
- monitoring of investments and exert of corporate governance after providing finance,
- facilitating trading of financial instruments, risk diversification, management of risk,
- mobilizing and pooling of savings,
- ease the exchange of goods and services.

Each of these financial functions may influence savings and investment decisions and hence economic growth".<sup>7</sup>

As the above fragment indicates financial development means, first of all, changes of qualitative character. These functions are implemented by financial markets and financial intermediaries. In the long run implementation of these functions leads to an increase in capital accumulation. Furthermore, through creating possibilities of risk diversification and creating finance sources, they stimulate increase in technological innovations. Together, they stimulate economic growth. J. Greenwood and B. Jovanovich indicated parallels and interdependencies in financial market development and economic growth. According to them, economic growth provides means thanks to which financial markets and financial intermediation develop. In turn, this process accelerates economic growth by supporting capital allocation.<sup>8</sup>

Moreover, the long-term relationships between the financial market development and economic growth is also worth mentioning. The financial market development is a component of a broadly understood financial development. Financial development is expressed, on the

<sup>&</sup>lt;sup>6</sup> See: Schumpeter (p.72-85, 155-202), Fiedor (1979, p. 21-30).

<sup>&</sup>lt;sup>7</sup> Levine (2004, p. 5-7)

<sup>&</sup>lt;sup>8</sup> See: Greenwood, Jovanovich (1989, p. 25).

one hand, by growing values of specific quantity variables characterising changes in the financial market and in the banking system which embrace, among others, an increase in the number of banks per 1,000 inhabitants, growth in bank assets in relation to GDP, an increase in the bank credit value in relation to GDP, an increase in stock exchange capitalisation in relation to GDP, an increase in the number of public partnerships (whose shares are listed at the stock exchange), an increase in the number of new emissions of financial instruments. On the other hand, one can speak about financial development when certain quality changes occur, like, e.g. launching new bank products and financial innovations, a tendency of economic units to invest savings in new financial products and use new financial services, emergence of new specialist financial institutions meeting new needs of the market, combining and permeating of hitherto separate types of financial activities and financial services (e.g. bank assurance), consolidation of financial institutions and their internationalisation.

According to R. Levine the influence of financial development on economic growth follows the block diagram presented in Figure 1.

It is worth mentioning that the correlations between financial development and, in particular, financial markets and economic growth, is not unambiguous from the theoretical point of view as well as from the point of view of empirical studies based on different econometric methods. J. Robinson claimed that "where enterprise leads, finance follows".<sup>9</sup> R. Lucas also questioned the relationship between financial development and economic growth, claiming that if there is any such relationship, the role of finance in economic growth is overestimated.<sup>10</sup>

<sup>&</sup>lt;sup>9</sup> See: Robinson(1952, p. 80). <sup>10</sup> See: Lukas (1988).





Source: (1997, p. 691).

In recent years there were several publications which were highly skeptical about the impact of financial development on economic growth. These include publications by, among others, P. Wachtel, M.J. Manning and P. L. Rousseau.<sup>11</sup>

<sup>&</sup>lt;sup>11</sup> See: Manning (2003), Wachtel (2003), Rousseau, Wachtel (2005).

### 2. Review of methods and results of selected empirical investigations

Majority of empirical research into relationships between financial development and economic growth is based on panel data and thus concerns not particular countries but their groups.

A basic econometric model used in the research into relationships between financial development and economic growth is R. Levine and R.G. King's version of the R. Barro's model of economic growth regression. This model assumes the following form:

# $Y_{it} = \alpha_0 + \alpha F_{it} + \beta X_{it} + u_{it},$

where  $Y_{it}$  is the growth rate of real GDP per capita in i-th country over the period t,  $F_{it}$  - is a financial development index in the i-th country over the period t (the ratio of the financial sector's liquid liabilities to GDP, the ratio of credit for the non-financial private sector to GDP, the ratio of credit for the non-financial private sector to total domestic credits, the ratio of domestic assets of deposit banks to domestic assets of the entire banking sector),  $X_{it}$  – the vector of basic, predetermined instrumental variables explaining economic growth in the i-th country over the period of t (the natural logarithm of initial GDP per capita, the natural logarithm of the scholarisation index – the ratio of children registered in secondary schools to the total number of children at the school age, the share of foreign trade turnover in GDP, the ratio of government consumption to GDP, the ratio of budget deficit to GDP). The model is estimated with the use of the double ordinary least square (2OLS) method. After GDP per capita has been replaced in the equation with the per capita capital growth rate, and next with the per capita efficiency growth rate, and the investment rate in GDP - the authors investigated the influence of financial development indicators on these quantities using the same model form. It is also worth mentioning that R.G. King and R. Levine used the statistical panel data based on the World Bank's statistics - Word Development Indicators and partly those of IFS – International Financial Statistics of the International Monetary Fund. On the other hand, P.L. Rousseau and P. Wachtel used the World Bank's database - World Development Indicators. R. King and R. Levine carried out research into relationships between the figures of real GDP per capita and the size of financial intermediation measured by the ratio of the financial system's liquid liabilities to GDP, based on a sample consisting of 80 countries and covering the 1960-1989 period. Next, they investigated the influence of financial development indicators on long-term economic growth rates per capita, capital accumulation and productivity growth. In each case the correlation indicators were high and statistically significant but different depending on a group of countries - divided into countries of low, medium and high economic development.<sup>12</sup> Further research based on panel data also confirmed a relatively strong impact of financial development, including that of financial markets, on economic growth.<sup>13</sup>

On the other hand, P.L. Rousseau and P. Wachtel's research did not fully confirm the results obtained by the above mentioned authors. P.L. Rousseau and P. Wachtel applied the same research method as R. Levine and R.G. King, and used in their research the annual panel data comprising statistical data from 84 countries for the 1960-2003 period.<sup>14</sup>

M. Neimke investigated the relationships among the selected financial development indicators (among others, the ratio of M3 to GDP, the ratio of stock-exchange capitalization to GDP, the ratio of bank credit for enterprises to GDP, the share of state bank assets in total bank assets) and GDP growth, and investments and productivity in the countries undergoing transformations (countries of Central and Eastern Europe and former Soviet republics in Asia). To this end he used the panel data for the period of 1990-2000 and constructed 18 equations. For their estimation he applied the Ordinary least square methods (OLS Method) and for some of the equations also a generalized method of moments (GMM). On the basis of empirical investigations he proved that also in transition economies (including Poland), there is a significant positive correlation between the financial market development (and, in particular, capital market) and investments, productivity and GDP per capita. However, it is worth mentioning that taking into account the magnitude of the coefficient of determination  $R^2$  (values ranging from 0.008 to 0.25 and in one case – 0.48), the results obtained to a very small degree explain relationships between financial development and economic growth.<sup>15</sup> What is more, the economies examined are remotely comparable. It is difficult to compare Central European economies with those of the former Soviet republics in Asia due to extreme differences in their financial structures, financial market development and levels of economic growth.

<sup>&</sup>lt;sup>12</sup> See: King and Levine (1993, pp. 717-737).

 <sup>&</sup>lt;sup>13</sup> See: Levine, et al. (2000, pp. 31-77); Caporale, et al.(2004, pp. 33-50); Caporale, et al.(2005, pp. 166-176).
<sup>14</sup> See: Wachtel (2003), Rousseau, Wachtel (2005)

<sup>&</sup>lt;sup>15</sup> Neimke (2003, pp. 2-32).

# 3. Results of empirical research into relationships between financial market development and economic growth in Poland and selected EU countries in the years 1994-2007

### 3.1 Statistical data and indicators

The research used statistical data comprised in the database *"Financial structure dataset* (*Nov. 2008*)" developed by T. Beck and E. Al.-Hussainy according to the methodology described in *"A New Database on Financial Development and Structure*" by T. Beck, A. Demiurguç – Kunt and R. Levine. The statistical data used in the research and concerning shaping of financial development indicators are derived entirely from this database. On the other hand, the data referring to GDP per capita and physical capital and investments are derived from the databases of Eurostat and the Central Statistical Office (GUS), as well as the World Development Indicator base of the World Bank.<sup>16</sup>

The data referring to Poland, Greece, Italy cover the years 1994-2007 (14 years) and for Ireland cover the years 1996-2007(12 years). In the case of Ireland, available date for years 1994-1995 are not existing. The author has chosen the above period for his analysis due to the fact that in Poland only since 1994 it has been possible to speak about significant financial development indicators in relation to GDP. Thus, the work covers 14 observations concerning the following quantities: the growth rate of real GDP per capita (*GDPp*), the growth rate of real gross physical capital (*CAPITAL*), efficiency growth rate (*EFF*), the growth rate of real gross investment per capita (*GINV*) and the five financial indicators presented below.

A relatively small sample entailed several problems. Due to the restricted number of degrees of freedom in an econometric model, the number of explanatory variables had to be restricted, which, obviously affected the quality of obtained models. Moreover, data inconsistency may have occurred due to the fact that not all needed quantities were available in statistics of one database. Particularly, the data referring to the growth rates of real gross investment per capita and physical capital per capita were taken from other sources than the World Development Indicators of the World Bank, and more specifically from the databases of EUROSTAT and GUS. For example, there is no data concerning physical capital and its growth for Poland and other new EU countries in the World Bank and Eurostat bases. Therefore, the author used here a variable known as "gross fixed assets", which is available in Statistical Yearbooks published by GUS.

<sup>&</sup>lt;sup>16</sup> See: Beck, Al.-Hussainy (2008); Beck, Demiurguç – Kunt, (1999).

In order to determine the efficiency growth indicator (*EFF*), combining in itself effects of technology use, human resource and labour productivity, the method applied by R.G. King and R. Levine was used.<sup>17</sup> The starting point is an economic growth equation in the form:

$$y = k^a x$$
,

where y stands for real GDP per capita, k represents real physical capital per capita, x stands for remaining determinants of per capita GDP growth (joint factor of technology, human resources and labour), a –is a parameter of production function. This equation can be transformed by logarithming the sides of the equation to the form:

$$ln y = alnk + lnx$$

By changing symbols, this dependence can be converted (as R.G. King and R.Levine did it) to the form: GDPp = aCAPITAL + EFF, where GDPp represents the growth rate of real GDP per capita, CAPITAL – stands for the growth rate of physical capital per capita, and EFFindicates the efficiency growth rate, a – is a share of physical capital growth in real per capita GDP growth. Hence, EFF can be determined in the following way:

EFF = GDPp - aCAPITAL. Following R.G. King and R.Levine, the value a = 0.3.<sup>18</sup> Additionally, the growth rate of real gross investment per capita (*GINV*) was used in the study.

Financial market development indicators were constructed in the following way: <sup>19</sup> STOCK – the ratio of the stock-exchange capitalization on the stock market to GDP  $\{0,5|STOCK_t/P_e_t + STOCK_{t-1}/P_e_{t-1}\}/GDP_t/P_a_t$ 

*PBONDS* – the ratio of the stock-exchange capitalization on the terasury bonds market to GDP

 $\{0,5[PBONDS_t/P_e_t + PBONDS_{t-1}/P_e_{t-1}]\}/GDP_t/P_a_t,$ 

In the above formulas the following symbols were adopted:  $P_e_t$ -stands for an inflation rate (CPI) at the end of the year,  $P_e_{t-1}$  is an inflation rate (CPI) at the beginning of the year,  $P_a_t$ -represents the average yearly inflation rate (CPI),  $GDP_t$  - GDP in the year t, t-year.

# 3.2 Research method

In the study an econometric multi-equation model (simple model) was used. The model consists of 3 equations, each of which is estimated separately by the OLS method. These equations were constructed in such a way as to show the effect of both the real sphere and the

<sup>&</sup>lt;sup>17</sup> King, (1993, p. 722).

<sup>&</sup>lt;sup>18</sup> Ibidem.

<sup>&</sup>lt;sup>19</sup>See: Beck, Demiurguç – Kunt, (1999); Beck, Al.-Hussainy (2007).

financial sphere on economic growth, or on its major factors: capital growth and efficiency growth. The choice of the model was dictated by the following premises:

- Basic exogenous and endogenous models of growth indicate the following economic growth factors: physical capital, human capital, labour, technology; for this reason the dependencies between per capita real GDP growth and per capita physical capital growth (*CAPITAL*) and efficiency growth (*EFF*) constituting a joint effect of labour, human capital and technology productivity growth were adopted.
- 2) Conclusions drawn from the financial development theory point out that individual components of financial development affect economic growth by creation of possibilities to accumulate capital (physical capital growth) and technological innovation (*EFF* growth), which lead to economic growth.
- 3) A tendency to examine relationships between the real and financial spheres of economy, and, in particular, between the indicators characterizing financial market development and economic growth and its factors – physical capital, efficiency and investment.
- A relatively short time sequence limiting the number of explanatory variables due to the required number of degrees of freedom as well as limiting possibilities of stylized fact identification.

The first equation characterizes the correlation between per capita real GDP growth and its main factors:

(1)  $l_GDPp = l_{a_{10}} + a_{11} l_EFF + a_{12} l_CAPITAL + u_1$ ,

where  $l\_GDPp$  – represents the logarithm of per capita GDP growth,  $l\_EFF$  – stands for the logarithm of efficiency growth rate,  $l\_CAPITAL$  – is the logarithm of the growth rate of capital per capita,  $l\_a_{10}$  – stands for the logarithm of the constant.

The second equation characterizes the correlation between the growth rate of real GDP per capita and efficiency growth (*EFF*) and the ratios of the stock market capitalization to GDP and the treasury bond-market capitalization to GDP (*PBONDS*).

(2)  $GDPp = a_{20} + a_{21}EFF + a_{22}STOCK + a_{23}PBONDS + u_2$ 

The third equation characterizes the relationship between the growth rate of real physical capital per capita (*CAPITAL*) and the growth rate of gross investment, and the ratios of the stock market capitalization to GDP (*STOCK*) and treasury bond market capitalization to GDP (*PBONDS*).

(3)  $CAPITAL = a_{30} + a_{31}GINV + a_{32}STOCK + a_{33}PBONDS + u_3$ 

# 3.3 Presentation of research results

The first equation reveals the obvious significant stochastic relationships between GDP per capita and explanatory (independent) variables being the economic growth factors. The relationship between per capita GDP growth and per capita physical capital growth and efficiency growth is obvious from the point of view of economic growth models (see: Tables 1A,B,C,D). Due to the non-linear dependence (power function), the author used logarithms of the per capita real GDP growth rate and logarithms of the growth rates of real physical capital per capita and efficiency.

Table 1A Poland. Equation 1: OLS Method estimation with the use of 14 observations1994-2007

		Depend	ant v	ariable:	l_GDPp			
	Coej	ficient	St	andard	t-Student	p v	value	
				error				
Const	-0.1	54094	0.2	227312	-0.6779	0.5	1184	
l_CAPITAL	0.20	0.204385		646612	3.1609	0.0	0906	***
1_EFF	2FF 0.65		0.0	214507	507 30.7169		< 0.00001	
Mean of dependent	f dependent -		1880 Stand		ndard deviation of		0.54	49790
variable				depe	endent variable			
Sum of squared resi	duals	0.04	1929	Stan	dard error of		0.00	51739
				resid	luals			
Unadjusted R-squar	ed	0.98	9330	Adjı	isted R- square	d	0.98	87390
F(2, 11)		509.	9452	p va	lue for the F-te	st	1.4	3e-11
Log - Likelihood		20.8	1067	Durl	oin-Watson stat	tistics	2.54	41708
Autocorrelation of		-0.30	3210					
Residuals - rho1								

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

Table. 1B Greece.	<b>Equation 1: OI</b>	LS Metho	od estim	ation with	the use of 14	<b>1</b> observations
		1994	-2007			
	_	-		~ ~ ~ ~		

**Dependant variable: l\_GDPp** 

	Coej	Coefficient		anda	rd	t-Student	<i>p</i> .	value	
				terror					
Const	0.1	3191	0.0632669		669	2.0850	0.06117		*
1_EFF	0.77	.771421		0.0217344		35.4930	<0.	00001	***
1_CAPITAL	0.18	189911		1103	391	17.2034	<0.	00001	***
Mean of dependent		-3.473		3576 Stand		dard deviation	of	0.34	0671
variable			deper		depe	ndent variable			
Sum of squared resi	duals	0.00	4326 Stan		Stan	dard error of		0.01	9832
					resid	uals			
Unadjusted R-squar	ed	0.99	7132		Adju	usted R- squared		0.99	6611
F(2, 11)		1912.4			p val	ue for the F-te	st	1.04	4e-14
Log - Likelihood		36.7	0934		Durb	bin-Watson statistics		1.97	3208

Autocorrelation of	-0.069453		
Residuals - rho1			

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

#### Table 1C Ireland. Equation 1: OLS Method estimation with the use of 14 observations 1994-2007 Dependent variables L CDPp

		Depena	ent va	ariadi	e: 1_	_GDPp			
	Coef	fficient	Stand	lard		t-Student	<i>p</i> 7	value	
			terro	r					
Const	0.14	0.142907		0.0692364		2.0640	0.0	6342	*
1_EFF	0.84	0.843363		0.0105352		80.0517	<0.0	00001	***
CAPITAL 0.121		21753	0.0	163836	6	7.4314	0.0	0001	***
Mean of dependent		-2.69	95873 Stand		lard deviation	0.34	14828		
variable					eper	ndent variable			
Sum of squared resi	duals	0.00	0.002611		and	lard error of		0.01	15405
				residuals					
Unadjusted R-squar	red	0.99	8311	A	Adjusted R- square			0.99	98004
F(2, 11)	F(2, 11)		1.224	p ·	p value for the F-test			5.6	5e-16
Log - Likelihood 4			4567	D	urbi	in-Watson sta	tistics	1.95	52762
Autocorrelation of		-0.05	0234						
Residuals - rho1									

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

# Table 1D Italy. Equation 1: OLS Method estimation with the use of 14 observations 1994-2007

# **Dependent variable:** l\_GDPp

	Coej	ficient	Sta	andar	d	t-Student	p va	lue	
			t	terror					
Const	0.2	0.22232		0.116282		1.9119	0.08819		*
1_EFF	0.7	0.717412		0.0116509		61.5759	<0.0	00001	***
1_CAPITAL	0.26	0.268867		29652	23	9.0673	<0.0	00001	***
Mean of dependent		-4.0714		S	Stand	dard deviation	of	0.42	27651
variable				d	leper	ndent variable			
Sum of squared resi	duals	0.003583		R	Residual standard e		error	0.01	19953
Unadjusted R-squar	red	0.99	8219	C	Corre	orrected R squared		0.99	97823
F(2, 11)		2521	1.968	р	val	ue for the F-te	st	4.2	5e-13
Log - Likelihood		31.67							
Autocorrelation of									
Residuals - rho1									

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

In the case of Poland Equation 2 indicates a statistically significant positive relationship between real GDP growth per capita and the ratio of stock-market capitalization to GDP. An increase in stock-market capitalization by 1 percentage point causes an increase in per capita GDP growth rate by 0.021 percentage point. On the other hand, the relationship between the ratio of treasury bond-market capitalization to GDP and the growth rate of GDP per capita is negative, hence an increase in the share of the bond-market capitalization in GDP by 1 percentage point causes a decrease in the growth rate of real GDP per capita by 0.051 percentage point (see: Table 2A). The influence of both variables on the growth rate of real GDP per capita is definitely lower than that of efficiency.

Table 2A Poland. Equation 2: OLS Method estimation with the use of 14 observations1994-2007

		Depen	dent	variable	: GDPp			
	Coej	ficient	Sta	andard	t-Student	<i>p</i>	value	
			1	value				
Const	0,01	59835	0,00	)161021	9,9263	<0,	00001	***
EFF	1,040		013 0,02		50,2914	<0,	<0,00001	
STOCK	ТОСК 0,02			057401	3,8273	0,0	0333	***
PBONDS -0,0		511506	0,00	)926045	-5,5236	0,0	0025	***
Mean of dependent	Mean of dependent		0,047357		ndard deviation	of	0,0	19069
variable				dep	endent variable	•		
Sum of squared rest	iduals	0,00	0019	Star	ndard error of		0,0	01361
				resi	duals			
Unadjusted R-square	red	0,99	6079	Adj	usted R- square	ed	0,9	94902
F(2, 11)		846,	7375	p va	lue for the F-te	est	2,5	51e-12
Log - Likelihood		74,8	7862	Dur	bin-Watson sta	tistics	3,2	209570
Autocorrelation of		-0,62	9110					
Residuals - rho1								

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

In the case of Greece the relationship between the ratio of stock-market capitalization to GDP and the growth rate of real GDP per capita is statistically insignificant, whereas the relationship between the ratio of treasury bond-market capitalization to GDP and the growth rate of real GDP per capita is statistically significant and positive. It must be noted, however, that the applied equation does not explain these relationships very well because the low value of the Durbin- Watson test points to autocorrelation of the random component (see also: Table 2B).

					<u> </u>			
	Coej	ficient	Sta	andard	t-Student	p v	value	
			t	error				
Const	-0.01	0.0115715 0		)499579	-2.3162	0.0	4305	**
EFF	1.1	1.14482		818234	13.9914	<0.0	00001	***
STOCK	0.00	0.00197107 (		)175339	1.1241	0.2	8721	
PBONDS	0.0172936		0.00	)690463	2.5046	0.0	3120	**
Mean of dependent	0.032		2643	Standard deviation of			0.01	10493
variable				dep	endent variable			
Sum of squared resi	duals	0.00	0044 Stand		dard error of		0.00	02088
				resi	duals			
Unadjusted R-squar	ed	0.96	9524	Adj	usted R- square	ed	0.96	50381
F(2, 11)	106.		0417	p va	value for the F-test		7.0	3e-08
Log - Likelihood	ood 68.8		8855	Dur	bin-Watson sta	tistics	0.59	94337
Autocorrelation of		0.83	6106					
Residuals - rho1								

# Table 2B Greece. Equation 2: OLS Method estimation with the use of 14 observations 1994-2007 Dependent variable: *GDPp*

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

Irish economy reveals a statistically significant positive relationship between the ratio of stock-market capitalization to GDP and the growth rate of real GDP per capita. However, the effect of this explanatory variable on the explained variable is weaker than in Poland. An increase in the ratio of stock-market capitalization to GDP by 1 percentage point causes an increase in the growth rate of GDP per capita by 0.0093 percentage point. The relationship between the ratio of treasury bond-market capitalization to GDP and the growth rate of real GDP per capita is negative, but statistically insignificant (see: Table 2C).

Table 2C Ireland. Equation 2: OLS Method estimation with the use of 12 observations1996-2007

		Depen	dent	variable	: GDPp			
	Coef	ficient	Sta	andard	t-Student	p v	value	
			error					
Const	0.005	0.00567257		)386259	1.4686	0.1	0.18013	
EFF	0.99	0.998107		)31766	31.4206	<0.0	00001	***
STOCK	0.009	931736	0.00	)469812	1.9832	0.0	8264	*
PBONDS	PBONDS -0.00		0.0	124469	-0.1163		1032	
Mean of dependent		0.070250		Stan	dard deviation	of	0.02	5641
variable				dependent variable		<b>e</b>		
Sum of squared resi	duals	0.000013		13 Standard error of			0.00	1284
				residuals				
Unadjusted R-squared		0.99	8177	Adjı	Adjusted R- squared		0.99	7493
F(2, 11)		1459	9.978	p va	p value for the F-test		2.72	2e-11
Log - Likelihood		65.3	0044	Dur	oin-Watson sta	ntistics	1.40	9016

Autocorrelation of	0.265218	
Residuals - rho1		

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

Italian economy reveals a positive, statistically significant relationship between the ratio of stock-market capitalization to GDP and the growth rate of real GDP per capita and between the treasury bond-market capitalization to GDP ratio and the growth rate of real GDP per capita. It is noteworthy, however, that the ratio of stock-market capitalization to GDP has a weaker impact on the growth rate of real GDP per capita in Italy than in Poland and Ireland.

# Table 2D Italy. Equation 2: OLS Method estimation with the use of 14 observations1994-2007

	Coej	ficient	Sta	andard	t-Student	p v	value	
			e	error				
const	-0.00	-0.0070465 0		)325218	-2.1667	0.0	5548	*
EFF	1.0	1.00648		256856	39.1847	<0.0	00001	***
STOCK	0.004	0.00419153 0		0151216	2.7719	0.0	1972	**
PBONDS	0.01	0.0108575 0		)337861	3.2136	0.0	0928	***
Mean of dependent		0.0162		Stan	tandard deviation of			09358
variable				depe	ndent variable			
Sum of squared resi	duals	5.47	7e-06 Stand		dard error of		0.0	00739
			resid		residuals			
Unadjusted R-squar	red	0.99	5199	Adju	ljusted R- squared		0.9	93759
F(2, 11)		690.	9587	p val	value for the F-test		6.8	9e-12
Log - Likelihood		83.42		Durt	oin-Watson sta	tistics	1.3	85148
Autocorrelation of		0.30	3317					
Residuals - rho1								

# **Dependent variable:** *GDPp*

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

Stock-market capitalization affects also the growth of real physical capital per capita. In the case of Poland this effect is higher than the influence of the growth rate of real gross investment (see: Table 3A). An increase in the ratio of stock-market capitalization to GDP by 1 percentage point increases the growth rate of real GDP per capita by 0.066 percentage point (0.027 percentage point in the case of the growth rate of real gross investment). A negative impact of the treasury bond-market capitalization to GDP ratio on the growth rate of real physical capital per capita is also significant (an increase by 1 percentage point causes a

decrease in the growth rate of GDP per capita by 0.153 percentage point).

	Coej	fficient	Sta	andard		t-Student	p v	value	
			e	error					
const	0.05	53085	0.00	0.00519872		10.2112	< 0.00001		***
GINV	0.02	.0275522		0.0128545		2.1434	0.05770		*
STOCK	0.06	0.0665638 ( -0.153156 (		0.0187168 0.0301973		3.5564	0.00521		***
PBONDS	-0.1					-5.0718	0.0	0048	***
Mean of dependent		0.0305		Sta	Standard deviation of			0.0	08354
variable				dependent		ent variable			
Sum of squared resi	duals	0.00	0194	0194 Standard e		rd error of		0.0	04407
				res	residuals				
Unadjusted R-squar	ed	0.78	5905	Ad	ljuste	ed R- square	d	0.7216	
F(2, 11)		12.2	3608	рv	value	e for the F-te	st	0.0	01104
Log - Likelihood		58.432		Du	ırbin	-Watson star	tistics	3.0	84796
Autocorrelation of		-0.56	9964						
Residuals - rho1									

# Table 3A Poland. Equation 3: OLS Method estimation with the use of 14 observations 1994-2007 Dependent variable: *CAPITAL*

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

# Table 3B Greece. Equation 3: OLS Method estimation with the use of 14 observations1994-2007

Dependent variable. CATITAL									
	Coefficient		Standard		rd	t-Student p		,alue	
			e	error	•				
const	-0.0369099		0.0190626		526	-1.9362	0.0	8159	*
GINV	-0.00735161		0.0546895		395	-0.1344	0.89573		
STOCK	0.01	04382	0.00	)6548	888	1.5939	0.14205		
PBONDS	0.06	93212	0.0	2535	515	2.7344	0.0	2103	**
Mean of dependent	0.022511 Stan			dard deviation	0.01	0825			
variable			depe	ndent variable					
Sum of squared residuals		0.000635		1	Standard error of			0.00	)7971
				] ]]	residuals				
Unadjusted R-squar	0.582904			Adjusted R- squared			0.45	57775	
F(2, 11)		4.658431		1	p val	ue for the F-te	0.027575		
Log - Likelihood	50.13788		]	Durbin-Watson statistics			0.82	27001	
Autocorrelation of		0.61	7020						
Residuals - rho1				1					

Dependent variable: CAPITAL

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

In the economy of Ireland, like in Poland, a positive and statistically significant relationship

between stock-market capitalization and the growth rate of real physical capital per capita is revealed. What is more, an increase in the stock-market capitalization to GDP ratio has a stronger impact on the growth of real physical capital per capita in Ireland than in Poland. Like in the case of Poland, the relationship between treasury bond-market capitalization and the growth rate of real GDP per capita is negative and statistically significant (see: Table 3C).

# Table 3C Ireland. Equation 3: OLS Method estimation with the use of 12 observations 1996-2007 Dependent variable: *CAPITAL*

	Coeffi	cient	Stand	lard	t-Student	p value	2
			teri	ror			
const	0.0225	5292	0.006	5179	3.6461	0.0065	3 ***
GINV	0.0850	)699	0.024	1618	3.5208	0.0078	4 ***
STOCK	0.0369	9639	0.0076	58274	4.8113	0.0013	4 ***
PBONDS	-0.070	1747	0.021	8971	-3.2048	0.0125	2 **
Mean of dependent		0.03	6066	Stand	lard deviation of	of	0.005147
variable				deper	ndent variable		
Sum of squared resid	duals	0.00	00057	Stand	lard error of		0.002681
				resid	uals		
Unadjusted R-square	ed	0.80	2723	Adju	sted R- squared		0.728744
F(2, 11)		10.8	85071	p val	ue for the F-test	t	0.003420
Log - Likelihood		56.4	6559	Durb	in-Watson stati	stics	2.180682
Autocorrelation of		-0.09	0415				
Residuals - rho1							

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

In the case of Italy, the relationship between the ratio of stock-market capitalization to GDP and the ratio of treasury bond-market capitalization and the growth rate of real physical capital per capita is positive and statistically significant.

# Table 3D Italy. Equation 3: OLS Method estimation with the use of 14 observations1994-2007

Dependent variable: CAPITAL										
	Coefficient		Standard		t-Stude	t-Student p v		ralue		
			t	error						
Const	-0.01	61529	0.0	120238	-1.343	34	0.2	0883		
GINV	0.04	14347	0.0	390344	1.136	51	0.2	8241		
STOCK	0.01	10963	0.00	)536525	5 2.068	32	0.0	6548	*	
PBONDS	0.02	81834	0.0	125623	2.243	35	0.0	4872	**	
Mean of dependent		0.01	4519	Sta	indard devi	ation o	of	0.00	3461	
variable				dej	pendent var	iable				

Sum of squared residuals	0.000054	Standard error of	0.002327
·····		residuals	
Unadjusted R-squared	0.652474	Adjusted R- squared	0.548216
F(2, 11)	6.258276	p value for the F-test	0.011559
Log - Likelihood	67.37760	Durbin-Watson statistics	1.371166
Autocorrelation of	0.301120		
Residuals - rho1			

\*\*\* the variable is significant at the significance level of 0.01, \*\* the variable is significant at the significance level of 0.05, \* the variable is significant at the significance level of 0.1 Source: the author's own calculations with the use of the GRETL programme.

It is interesting that in the case of Greece and Italy a high ratio of public debt to GDP (above 100%) and a positive relationship between the treasury bond-market capitalization to GDP ratio and economic growth are indicated. However, in the case of Poland and Ireland the ratio of public debt to GDP is relatively low (below 60%) and the relationship between the bond-market capitalization to GDP radio and economic growth is negative. This phenomenon requires a thorough investigation on a bigger sample and with the use of panel data comprising two groups of countries: countries of high public debt with reference to GDP and countries of a relatively low ratio of public debt to GDP.

# Conclusions

An analysis of relationships between selected indicators of financial market development and economic growth in Poland in the period from 1994 to 2007 points to the following stylized facts:

- In the analyzed period, there was a statistically significant relationship between financial market development and economic growth in Poland, Greece, Ireland and Italy;
- An increase in the treasury bond-market capitalization (result of increased budget deficit and public debt) has a negative and relatively strong impact on the rate of real economic growth and the rate of real physical capital growth in Poland and Ireland, hence it may confirm a negative effect "portfolio crowding out" caused by increased budget deficit and public debt on economic growth; in the case of Greece and Italy, i.e. the countries of a high ratio of public debt to GDP, there is a statistically significant, positive relationship between the treasury bond-market capitalization to GDP ratio and economic growth;
- There is a statistically significant but relatively weak, positive relationship between the stock-market capitalization and growth rates of real GDP per capita and real

physical capital per capita; this relationship is stronger in the case of Poland than in the case of other examined countries.

To sum up, the conducted analysis revealed a statistically significant and meaningful impact of financial market development on economic growth in Poland, Ireland and Italy, but weaker in the case of Greece. The research done is of preliminary character and the conclusions formulated above require further verification by extensive studies covering a larger number of explanatory variables and a larger number of observations.

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